

Beyond global sensitivity:

Def: The local sensitivity of a query $f: \mathcal{D} \rightarrow \mathbb{R}$ at database D is

$$\Delta_{\text{LS}}(f, D) = \max_{D' \sim D} |f(D) - f(D')|$$

- would like to claim can just replace global sensitivity w/ LS.

- Unfortunately can't do this: magnitude of noise might leak info!

Ex: $f = \text{median}$

$$D = (\underbrace{0, 0, \dots, 0}_{\frac{n}{2} + 1}, 10^6, 10^6, \dots, 10^6)$$

$$D' = (\underbrace{0, 0, \dots, 0}_{\frac{n}{2}}, 10^6, 10^6, \dots)$$

$$f(D) = f(D') = 0$$

$$\Delta_{\text{LS}}(f, D) = 0 \quad \Delta_{\text{LS}}(f, D') = 10^6$$

so adding $L_2(\frac{\Delta_{\text{LS}}(f, D)}{\epsilon})$ noise definitely not DP:
on D always returns 0, on D' returns large values.

Can we still save this idea?
Basic idea: computing Δ_{CS} looks at database! need to privatize.

Propose-Test-Release:

- Idea: - Propose bound on LS
- Test if LS below bound (privately)
- If yes, answer $w(L)$ with
no, answer \perp (or use global sensitivity ϵ).

Formally:

- 1) Propose bound β on LS
- 2) compute $\delta = \text{distance from } D \text{ to } D' \text{ s.t. } \Delta_{CS}(f, D') \geq \beta$
(where distance is # points that change). $\Delta \delta = 1$!
- 3) compute $\hat{\delta} = \delta + \underbrace{L_{CP}(\frac{1}{\epsilon})}_2$
- 4) If $\hat{\delta} \leq \frac{1}{\epsilon} \ln(1/\delta)$ return \perp
- 5) else ($\hat{\delta} > \frac{1}{\epsilon} \ln(1/\delta)$) return $f(D) + L_{CP}(\frac{\beta}{\epsilon})$

Note: Step 2 can be computationally difficult! But
often easy, e.g., median

Thm: PTR is $(2\epsilon, \delta)$ -DP.

- Indication: - $\Delta \delta = 1$, so $\hat{\delta}$ PP by Laplace Mechanism
- Step 5 uses bound on LS that's good for both D, D'

Pr: Let $D \sim D' \in \mathcal{D}$

$$\Delta \hat{\gamma} = 1, \text{ so } \Pr[M(D) = \perp] \leq e^\epsilon \Pr[M(D') = \perp]$$

Two cases:

Case 1: $\Delta_{CS}(f, D) > \beta$

$$\Rightarrow \gamma = 0$$

$$\text{Let } S \subseteq \mathbb{R} \cup \{\perp\}$$

$$\Pr[M(D) \in S] = \Pr[M(D) \in S \cap \{\perp\}] + \Pr[M(D) \in S \cap \mathbb{R}]$$

$$\leq e^\epsilon \Pr[M(D') \in S \cap \{\perp\}] + \Pr[M(D) \in \mathbb{R}]$$

$$\leq e^\epsilon \Pr[M(D') \in S] + \frac{\Pr[Z \geq \frac{1}{\epsilon} \ln \frac{1}{\delta}]}{Z \sim \text{Lap}(\epsilon/\delta)}$$

$$\leq e^\epsilon \Pr[M(D') \in S] + \delta \quad \checkmark$$

Case 2: $\Delta_{CS}(f, D) \leq \beta$

\Rightarrow basic composition of 2 mechanisms:

- Laplace for set $\hat{\gamma}: (\epsilon, 0)$ -DP
- Laplace for return $f(D) + \text{Lap}(\beta/\epsilon)$

$$\Rightarrow \beta \geq \Delta_{CS}(f, D) \text{ by case 1}$$

$$\Rightarrow \epsilon\text{-DP}$$

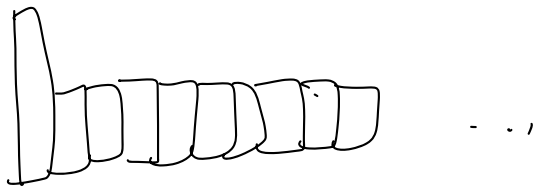
$$\Rightarrow 2\epsilon\text{-DP}$$

$$\Rightarrow (2\epsilon, \delta)\text{-DP} \text{ overall.}$$

Ex: Mode

sp want to return element that appears most often in D.

- Obvious approach: histogram, return highest count.
- Can add $\log(1/\epsilon)$ -noise to each count (histogram)
- sps $1 \leq k$, 1 element largest, all other el's 2nd largest:



\Rightarrow in expectation, same element gets noise

$$\approx \frac{\log k}{\epsilon}$$

\Rightarrow only works if gap b/w largest, 2nd largest $\geq \frac{\log k}{\epsilon}$ k can be huge!

Try using PTR instead.

- Repose $\beta = 0$
- Let $\gamma = \text{diff. b/w largest and 2nd-largest count}$
- Let $\hat{\gamma} = \gamma + \text{Lap}(1/\epsilon)$
- If $\hat{\gamma} \leq \frac{1}{\epsilon} \ln(1/\delta)$, return \perp
- else return most frequent el't

Thm: (ϵ, δ) -DP

PF: PTR w/ $\beta = 0$!

Thm: Returns the node w/ prob $\geq 1-\delta$ as long as diff b/w largest and 2nd largest count $\geq \frac{1}{\epsilon} \ln \frac{1}{\delta}$

PF: Only fails if $\hat{y} \leq \frac{1}{\epsilon} \ln \frac{1}{\delta}$

$$\Rightarrow -\text{Lap}\left(\frac{1}{\epsilon}\right) \geq \hat{y} - \frac{1}{\epsilon} \ln \frac{1}{\delta} \geq \frac{1}{\epsilon} \ln \frac{1}{\delta}$$

$$\Pr[-\text{Lap}\left(\frac{1}{\epsilon}\right) \geq \frac{1}{\epsilon} \ln \frac{1}{\delta}] \leq e^{-\ln \frac{1}{\delta}} = \delta \quad \checkmark$$

Generalize: "stable histograms".

Laplace histograms: some entry off by $\Theta\left(\frac{\log k}{\epsilon}\right)$

k could be huge: arbitrarily larger than n ! (??)
Even infinite?

Want to get $\Theta\left(\frac{1}{\epsilon} \ln \frac{1}{\delta}\right)$.

Think of $\delta \leq \frac{1}{n}$, ^{w/ 101} so looking for $\frac{1}{\epsilon} \ln n$ error.

Idea: could be done if only released nonzero counts, but would violate privacy

- only release counts above a threshold, round all others down to 0!

SH:

- For each $x \in X$:

- if $D_x = 0$, set $a_x = 0$

($D_x = 0$ not a fix 0)

- if $D_x > 0$:

- set $a_x = D_x + C_p(\frac{1}{\epsilon})$

- if $a_x < \frac{1}{\epsilon} \ln \frac{1}{\delta} + 1$, set $a_x = 0$

Thm: For $\delta \leq \frac{1}{n}$, w.h.p. $|a_x - D_x| \leq O(\frac{1}{\epsilon} \ln \frac{1}{\delta}) \quad \forall x \in X$

PF: Definitely true if $D_x = 0$

else, w.h.p. $D_x + C_p(\frac{1}{\epsilon}) \leq D_x + O(\frac{1}{\epsilon} \ln n)$

$\leq D_x + O(\frac{1}{\epsilon} \ln \frac{1}{\delta}) \quad \forall x \in X$

$\Rightarrow |a_x - D_x| \leq O(\frac{1}{\epsilon} \ln \frac{1}{\delta}) + \frac{1}{\epsilon} \ln \frac{1}{\delta} + 1$

$\leq O(\frac{1}{\epsilon} \ln \frac{1}{\delta}) \quad \text{w.h.p.}$

Thm: (ϵ, δ) -DP.

PF: Let $D \sim D'$ where $D_x = D'_x + 1$ (add 1 to D)

\Rightarrow only diff. b/w $SH(D)$ and $SH(D')$ is in

a_x (all other values identically distributed)

$$\underline{\text{Case 1: } p'_x > 0}$$

\Rightarrow add $\mathcal{L}_{\text{CP}}(\frac{1}{\epsilon})$ -noise to both p'_x and p_x

$\Rightarrow \epsilon$ -DP (function is post-processing)

$$\underline{\text{Case 2: } p'_x = 0 \Rightarrow a'_x = 0}$$

$$\Rightarrow p_x = 1$$

$$\Rightarrow \Pr[a_x \neq 0] = \Pr[\mathcal{L}_{\text{CP}}(\frac{1}{\epsilon}) > \frac{1}{\epsilon} \ln \frac{1}{\delta}] \leq \delta$$

$$\Rightarrow (0, \delta)\text{-DP}$$

$$\Rightarrow (\epsilon, \delta)\text{-DP overall.}$$