

# Lecture 13

## Two Examples of Bayes-Kalman.

Note Title

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Recall.

$$Y_t = \{y_1, \dots, y_t\}, \quad \text{state } X_t$$

prediction

$$P(X_{t+1} | Y_t) = \sum_{X_t} P(X_{t+1} | X_t) P(X_t | Y_t)$$

correction

$$P(X_{t+1} | Y_{t+1}) = \frac{P(y_{t+1} | X_{t+1}) P(X_{t+1} | Y_t)}{P(y_{t+1} | Y_t)} \quad \frac{P(X_{t+1} | X_t)}{P(y_t | X_t)}$$

Particle filters:

Approximate  $P(X_t | Y_t)$  by samples

$(X_t^1, \dots, X_t^m)$  define recursively

update by sampling from  $P(X_{t+1} | X_t^i)$  to get  $\{X_{t+1}^1, \dots, X_{t+1}^m\}$  represents  $P(X_{t+1} | Y_t)$ :

give each sample a weight  $w_{t+1}^i \propto P(y_{t+1} | X_{t+1}^i)$   
 ("∝" means proportional to)

Resample from  $\{X_{t+1}^1, \dots, X_{t+1}^m\}$

with probability proportional to  $w_{t+1}^i$  (with replacement) to get new samples

$\{X_{t+1}^{1*}, \dots, X_{t+1}^{m*}\}$  represent  $P(X_{t+1} | Y_{t+1})$

## (2) Example 1: Tracking. Gaussian Models.

state variable  $x_t = (x_{t,1}, x_{t,2})$

$x_{t,1}$  location  $x_{t,2}$  velocity

$$\text{Update: } x_{t,1} = x_{t-1,1} + x_{t-1,2} + \frac{1}{2} \omega_t$$

$$x_{t,2} = x_{t-1,2} + \omega_t$$

$\omega_t$  i.i.d. normal  $N(0, q^2)$ .

Equivalently:  $P(x_t | x_{t-1}) = P(x_{t,1} | x_{t-1,1}, x_{t-1,2}) P(x_{t,2} | x_{t-1,2})$ .

$$\text{with } P(x_{t,1} | x_{t-1,1}, x_{t-1,2}) = N(x_{t-1,1} + x_{t-1,2}, q^2/4)$$

$$P(x_{t,2} | x_{t-1,2}) = N(x_{t-1,2}, q^2)$$

Observations  $y_t = x_{t,1} + v_t$   $v_t \sim N(0, r^2)$

Note: we only observe the positions  $x_{t,1}$ . We cannot directly observe the velocity  $x_{t,2}$ .

$$P(y_t | x_t) = N(x_{t,1}, r^2)$$

For this example all the distributions are Gaussians. So Bayes-Kelmon reduces to updating means and covariances.

Next we consider a harder problem that needs particle filters. The target object may be invisible and there may be distractor objects present.

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Example 2: The target object has a Gaussian prior  $P(x_{t+1}|x_t) = N(x_{t+1}|\mu, \sigma_p^2)$  and the distribution of the "true" observation is Gaussian  $N(x_t, r^2)$  if target is visible.

### The Observation Model

Observations occur in a window size  $\Delta$ .

At time  $t$ , the number of observations in the window is a random variable  $m_t$ , which can be measured.

The observations are represented by a vector:

$\underline{y}_t = (y_{t,1}, \dots, y_{t,m_t})$ , where  $y_{t,i}$  is position of  $i^{\text{th}}$  obs.

We assume that:

(1) the probability that target is visible is  $P_d$ ,

(2) the prob dist of number of distractors is Poisson, with parameter  $\lambda\Delta$ ,

(3) the position of a distractor is uniformly distributed in the window (i.e. prob density is  $1/\Delta$ )

(4) the position of target object, if present, is modeled by  $N(x_t, r^2)$

Define an indicator variable:

$$I_k = \begin{cases} 0, & \text{if target is invisible} \\ k, & \text{if target is } k^{\text{th}} \text{ observation.} \end{cases}$$

$$P(\underline{y}_t | x_t, I_t=0) = \Delta^{-m_t} \frac{(\lambda\Delta)^{m_t}}{m_t!} e^{-\lambda\Delta} = \frac{\lambda^{m_t}}{m_t!} e^{-\lambda\Delta}$$

$$P(\underline{y}_t | x_k, I_t=k) = \frac{\lambda^{m_t-1}}{(m_t-1)!} e^{-\lambda\Delta} \frac{1}{\sqrt{2\pi}r} \exp\left\{-\frac{(y_{t,k}-x_t)^2}{2r^2}\right\}$$

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By assumption (1):

$$P(I_t=0) = 1 - p_d, \quad P(I_t=k) = p_d/m_t.$$

Hence:  $P(\underline{y}_t, I_t | X_t) = P(\underline{y}_t | I_t, X_t) P(I_t)$

(Note  $\alpha$  is proportional)  $\propto (1 - p_d) \lambda$ , if  $I_t = 0$   
 $p_d \frac{1}{(2\pi)^{1/2} r} \exp\left\{-\frac{(y_{t,k} - x_t)^2}{2r^2}\right\}$ , if  $I_t = k$ .

where the proportionality factor is  $\frac{\lambda^{m_t-1} e^{-\lambda \Delta}}{m_t!}$ .

Now,  $I_t$  is unknown (it is a missing/latent/hidden variable) so we must sum it out to obtain the observation model:

$$P(\underline{y}_t | X_t) = \sum_{I_t} P(\underline{y}_t, I_t | X_t) \propto (1 - p_d) \lambda + \sum_{k=1}^{m_t} p_d \frac{1}{(2\pi)^{1/2} r} \exp\left\{-\frac{(y_{t,k} - x_t)^2}{2r^2}\right\}$$

$\propto$   
proportional to.

The observation model is non-Gaussian so we cannot apply Kalman's update equations

Instead we can use particle filters. This only requires:

(i) that we can draw samples from  $p(x_{t+1} | x_t)$

— easy because it is Gaussian

(ii) that we can evaluate  $P(\underline{y}_t | X_t)$  up to a normalization constant. — which is also easy in this case.

Note that we can use particle filtering for this problem even if  $p(x_{t+1} | x_t)$  is non-Gaussian, provided we can sample from it, and even if the likelihood term  $p(\underline{y}_t | X_t)$  is more complicated.